

## On the Ratio of Exponentiated Exponential Hypergeometric Function and Gamma Random Variables

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### Abstract

The present paper deals with the distribution of ratio  $Z = X/Y$  when  $X$  and  $Y$  are independently distributed as three parameter exponential distribution involving Hypergeometric function and two parameter gamma random variables respectively. The p.d.f. ,c.d.f. , moments and cumulants are also derived for the distribution

**Keywords:** Exponentiated exponential distribution, Hypergeometric function, Reliability, two parameter gamma distribution, moments, cumulants.

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### Introduction

Joshi & Modi [5] introduced exponentiated exponential distribution with three parameters and its unique form with Hypergeometric function was given by Khan [6]. The ratio  $Z = X/Y$  has vast application in real life and it is extensively studied by many researchers like Joshi & Joshi [4], Marsaglina [7], Nadarajah [8]. In the present paper we derive distribution of  $X/Y$  when  $X$  and  $Y$  are independent random variables having exponentiated exponential distribution involving Hypergeometric distribution.

$$f_x(x) = \frac{\lambda\alpha\beta e^{-\lambda x} {}_2F_1(1-\alpha, \beta_1; \gamma; \beta e^{-\lambda x})}{1 - {}_2F_1(-\alpha, \beta_1; \gamma; \beta)} \quad (1.1)$$

Where

$$\gamma > 0, |\beta e^{-\lambda x}| < 1, x > 0, \alpha > 0, \lambda > 0, |\beta| < 1.$$

and two parameter gamma function with p.d.f.given as-

$$f_Y (y) = (a^m / \Gamma m) e^{-ay} y^{m-1}; \quad (1.2)$$

for

$$m > 0, a > 0, y > 0.$$

### P.d.f. and c.d.f. Ratio of X & Y.

**Theorem 2.1:** If X and Y are distributed according to equations (1.1) and (1.2) respectively then c.d.f. and p.d.f. of  $Z=X/Y$  can be given as -

$$F_z(z) = \frac{\alpha\beta}{1-{}_2F_1(-\alpha, \beta_1; \gamma; \beta)} \sum_{n=0}^{\infty} \frac{(1-\alpha)_n (\beta_1)_n \beta^n (a^m - [\lambda z(1+n) + a]^m)}{(\gamma)_n (n+1)! [\lambda z(1+n) + a]^m} \quad (2.1)$$

and

$$f_z(z) = \frac{\lambda\alpha\beta}{1-{}_2F_1(-\alpha, \beta_1; \gamma; \beta)} \sum_{n=0}^{\infty} \frac{(1-\alpha)_n (\beta_1)_n \beta^n a^m (-m)}{(\gamma)_n (n)! [\lambda z(1+n) + a]^{m+1}} \quad (2.2)$$

where;  $x > 0, \alpha > 0, \beta_1 > 0, \lambda > 0, 0 < \beta < 1$ .

**Proof:** Lets  $X/Y = Z$  then its c.d.f. can be defined as

$$\begin{aligned} F_z(z) &= P\left(\frac{X}{Y} \leq Z\right) = P(X \leq YZ) \\ &= \int_0^{\infty} F_X(yz) f_Y(y) dy \\ &= (a^m / \Gamma m) \int_0^{\infty} \frac{\alpha\beta}{1-{}_2F_1(-\alpha, \beta_1; \gamma; \beta)} \sum_{n=0}^{\infty} \frac{(1-\alpha)_n (\beta_1)_n \beta^n}{(\gamma)_n (n+1)!} (e^{-\lambda yz(1+n)} - 1) e^{-ay} y^{m-1} dy. \end{aligned}$$

where  $F_X(yz)$  is obtained by using (Eq.(1.2) page 252, khan[6] )and finally by using well known formula for gamma function

$$\int_0^{\infty} e^{-ax} x^{n-1} dx = \frac{\Gamma n}{a^n}; n > 0$$

$$F_z(z) = \frac{\alpha\beta}{1-{}_2F_1(-\alpha, \beta_1; \gamma; \beta)} \sum_{n=0}^{\infty} \frac{(1-\alpha)_n (\beta_1)_n \beta^n (a^m - [\lambda z(1+n) + a]^m)}{(\gamma)_n (n+1)! [\lambda z(1+n) + a]^m}$$

and differentiation above equation w.r.t z we can easily obtained

$$f_z(z) = \frac{\lambda\alpha\beta}{1-{}_2F_1(-\alpha, \beta_1; \gamma; \beta)} \sum_{n=0}^{\infty} \frac{(1-\alpha)_n (\beta_1)_n \beta^n a^m (-m)}{(\gamma)_n (n)! [\lambda z(1+n) + a]^{m+1}}$$

**Hazard Rate Function**

The hazard rate function defined by  $h(x) = \frac{f(x)}{1-F(x)}$  and for the variate  $z$  it is given as

$$h_z(z) = \frac{\frac{\lambda\alpha\beta}{1-{}_2F_1(-\alpha, \beta_1; \gamma; \beta)} \sum_{n=0}^{\infty} \frac{(1-\alpha)_n (\beta_1)_n}{(\gamma)_n (n)!} \frac{\beta^n a^m (-m)}{[\lambda z(1+n) + a]^{m+1}}}{1 - \frac{\alpha\beta}{1-{}_2F_1(-\alpha, \beta_1; \gamma; \beta)} \sum_{n=0}^{\infty} \frac{(1-\alpha)_n (\beta_1)_n}{(\gamma)_n (n+1)!} \frac{\beta^n (a^m - [\lambda z(1+n) + a]^m)}{[\lambda z(1+n) + a]^m}} \tag{3.1}$$

Further

$$h_z(0) = \frac{\lambda\alpha\beta (-m)_2 {}_2F_1(1-\alpha, \beta_1; \gamma; \beta)}{a [1 - {}_2F_1(-\alpha, \beta_1; \gamma; \beta)]} \tag{3.2}$$

and  $h(\infty) = 0$ , the hazard rate function ranges from  $h(0)$  to zero.

The Survival or Reliability function for c.d.f. of equation (1.2) is given as:

$$S(z) = 1 - F_z(z)$$

$$S(z) = 1 - \frac{\alpha\beta}{1-{}_2F_1(-\alpha, \beta_1; \gamma; \beta)} \sum_{n=0}^{\infty} \frac{(1-\alpha)_n (\beta_1)_n}{(\gamma)_n (n+1)!} \frac{\beta^n (a^m - [\lambda z(1+n) + a]^m)}{[\lambda z(1+n) + a]^m} \dots\dots\dots \tag{3.3}$$

**Moments**

The  $r$ th moments for random variable  $Z=X/Y$  whose p.d.f. is given by equation (2.2) is -

$$\begin{aligned} E(z^r) &= \frac{\lambda\alpha\beta}{1-{}_2F_1(-\alpha, \beta_1; \gamma; \beta)} \sum_{n=0}^{\infty} \frac{(1-\alpha)_n (\beta_1)_n}{(\gamma)_n (n)!} \frac{\beta^n a^m (-m)}{\int_0^{\infty} [\lambda z(1+n) + a]^{-m-1} z^r dz} \\ &= \frac{\alpha\beta}{1-{}_2F_1(-\alpha, \beta_1; \gamma; \beta)} \sum_{n=0}^{\infty} \frac{(1-\alpha)_n (\beta_1)_n}{(\gamma)_n (n)!} \frac{\beta^n (-m)}{\frac{a^r}{\lambda^r (1+n)^{r+1}} \mathbf{B}(m-r, r+1)} \end{aligned} \tag{4.1}$$

Where

$$B(p, q) = \frac{\Gamma p \Gamma q}{\Gamma p + q}$$

On taking  $r=1, 2, 3, 4$ . ( $m > r$ ), we can easily obtained first, second, third & fourth moments about origin.

### Relation between Cumulants and Moments

Since cumulants is defined as -

$$K(t) = \log_e M(t) = k_1 t + k_2 \frac{t^2}{2!} + k_3 \frac{t^3}{3!} + \dots + k_r \frac{t^r}{r!} + \dots \quad (5.1)$$

The coefficients  $k_1, k_2, \dots, k_r$  are called the first, second, .... rth cumulants. Further relation between moments and cumulants is given as.

$$\mu_r^\lambda = \sum_{j=1}^r {}^{r-1}C_{j-1} \mu_{r-j}^\lambda K_j. \quad (5.2)$$

By using (4.1), we have

$$\begin{aligned} & \frac{\Gamma(r+1)}{\lambda^r} \sum_{n=0}^{\infty} \frac{(1-\alpha)_n (\beta_1)_n}{(\gamma)_n (n)!} \frac{\beta^{n+1}}{(n+1)^{r+1}} \\ & \frac{\alpha\beta}{1-{}_2F_1(-\alpha, \beta_1; \gamma; \beta)} \sum_{n=0}^{\infty} \frac{(1-\alpha)_n (\beta_1)_n}{(\gamma)_n (n)!} \frac{\beta^n (-m)}{\lambda^r (1+n)^{r+1}} \mathbf{B}(m-r, r+1) \\ & = \sum_{j=1}^r {}^{r-1}C_{j-1} \frac{\alpha\beta}{1-{}_2F_1(-\alpha, \beta_1; \gamma; \beta)} \sum_{n=0}^{\infty} \frac{(1-\alpha)_n (\beta_1)_n}{(\gamma)_n (n)!} \frac{\beta^n (-m)}{\lambda^{r-j} (1+n)^{r-j+1}} \mathbf{B}(m-r+j, r-j+1) K_j \end{aligned} \quad (5.3)$$

On specializing  $j=1, 2, \dots, r$ , we can get the values of  $k_1, k_2, \dots, k_r$ .

### Particular Cases

**Case I:** Setting  $\beta_1 = \gamma$ ,  $a=1$ ,  $\lambda = 1$ , the p.d.f. & c.d.f. of  $z$  is given by

$$f_z(z) = \frac{\alpha\beta}{1-(1-\beta)^\alpha} \sum_{n=0}^{\infty} \frac{(1-\alpha)_n \beta^n (-m)}{n!} [z(1+n)+1]^{-m-1} \quad (6.1)$$

$$F_z(z) = \frac{\alpha\beta}{1-(1-\beta)^\alpha} \sum_{n=0}^{\infty} \frac{(1-\alpha)_n \beta^n}{(n+1)!} \left\{ \frac{1-[z(1+n)+1]^m}{[z(1+n)+1]^m} \right\} \quad (6.2)$$

**Case II:** Setting  $\beta = 1$ ,  $n=0$  and  $\alpha = -1$  in above equations we get

$$f_z(z) = \frac{m}{(z+1)^{m+1}} \quad (6.3)$$

$$F_z(z) = 1 - \frac{1}{(z+1)^m} \quad (6.4)$$

This is known result given by Joshi and Joshi [4, eq. (5.3) & (5.4) page 73]

**Case III:** Setting  $m=1$  in (6.3) & (6.4)

$$f_z(z) = \frac{1}{(z+1)^2} \quad (6.5)$$

$$F_z(z) = \frac{z}{(z+1)} \quad (6.6)$$

Which is also a known result given by Joshi and Joshi [4, eq. (5.5) & (5.6) page 74]

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